Dependence

Copula

Applications

1. What is the joint probability of concurrent heavy precipitation and high streamflow?

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What is the joint probability of low soil moisture and heatwave?

What is the joint probability of heatwave, drought severity and duration?



Precipitation

$$p_p = P(X \leq x)$$

Soil moisture

$$p_{sm} = P(Y \leq y)$$

$$p_{p-sm} = P(X \leq x, Y \leq y)$$



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Measures of linear dependence

$$s_{xy} = \frac{1}{n-1} \sum_{i=1}^{n} (x_i - \bar{x})(y_i - \bar{y})$$

$$r_{xy} = \frac{s_{xy}}{s_x s_y} = \frac{\sum_{i=1}^n (x_i - \bar{x})(y_i - \bar{y})}{\sqrt{\sum_{i=1}^n (x_i - \bar{x})^2} \sqrt{\sum_{i=1}^n (y_i - \bar{y})^2}}$$

Correlation Coefficient

 $-1 \le r_{xy} \le +1$

Correlation and Covariance



Correlation and Covariance



?

Correlation and Covariance







What is the dependence between large values of X and Y?

How large values are different than small values?

Where in the distribution there is a stronger relationship?





Large values of X and Y are strongly associated with each other

The dependence between large values is stronger than small values





The dependence between small values is stronger than large values

Transformation to uniform marginals

$$(x_{i}, y_{i}) \quad i = 1, ..., n$$

$$\left(\frac{n - R(x_{i}) + \frac{1}{2}}{n}, \frac{n - R(y_{i}) + \frac{1}{2}}{n}\right) \quad i = 1, ..., n$$

$$R(x_{i}) = \text{the rank of } x_{i} \text{ in the set } \{x_{1}, ..., x_{n}\}$$

Rank Correlation

The above transformation dissociates the correlation structure between variables from their marginal distributions.

Rank Correlation Methods:



Kendall's Rank Correlation Coefficient τ

Rank correlation methods measure the degree of monotone (increasing or decreasing) dependence (or association) between two variables.

$$\rho_s = 1 - \frac{6\sum_{i=1}^n d_i^2}{n(n^2 - 1)}$$

Here, d_i denotes differences between the ranks of two variables

X	Y	Rank (X _i)	Rank (Y _i)	d _i
9	28.4	1	1	0
15	29.3	2	2	0
24	37.6	3	7	-4
30	36.2	4	4.5	-0.5
38	36.5	5	6	-1
46	35.3	6	3	3
53	36.2	7	4.5	2.5
60	44.1	8	8	0
64	44.8	9	9	0
76	47.2	10	10	0
				Σd ² =32.5

$$\rho_s = 1 - \frac{6\sum_{i=1}^n d_i^2}{n(n^2 - 1)}$$

$$\rho_s = 1 - \frac{6(32.5)}{10(99)}$$

= 0.80

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For
$$n > 10$$

Z= $\rho_s \sqrt{n-1}$
= 0.80 * $\sqrt{9}$
=**2.4**

p-value = 1- Ø(z) = 1- normcdf(2.4) =**0.0082**

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strong positive dependence

Spearman's rho is the linear correlation between $F_1(X)$ and $F_2(Y)$, which are integral transforms of X and Y. In this sense it is a measure of rank correlation. Both $\rho_S(X,Y)$ and $\rho_\tau(X,Y)$ are measures of monotonic dependence between (X,Y). Both measures are based on the concept of **concordance**, which refers to the property that large values of one random variable are associated with large values of another, whereas discordance refers to large values of one being associated with small values of the other. • A copula is a multivariate probability distribution for which the marginal probability distribution of each variable is uniform. Copulas are used to describe the dependence between random variables.

$$p = P(X \le x, Y \le y)$$
$$p = C[F(X), G(Y)]$$

C is the copula and F(X) and G(Y) are the marginal cumulative distribution functions of precipitation (X) and soil moisture (Y), respectively

• Sklar's Theorem states that any multivariate joint distribution can be written in terms of univariate marginal distribution functions and a copula which describes the dependence structure between the variables.

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• Sklar's Theorem states that any multivariate joint distribution can be written in terms of univariate marginal distribution functions and a copula which describes the dependence structure between the variables.

$$p = P(X \le x, Y \le y)$$
$$p = C[F(X), G(Y)]$$

1) Choice of arbitrary marginal distributions:

They could take different forms; They could involve covariates.

2) Choice of an arbitrary copula function (dependence structure).

Copulas and Dependence



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Soil moisture

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Copulas and Dependence

• Copulas are popular in high-dimensional statistical applications as they allow one to easily model and estimate the distribution of random vectors by estimating marginals and copulae separately. • There are many parametric copula families available, which usually have parameters that control the strength of dependence.

Copula type	Function $C(u_1, u_2)$
Product	$u_1 u_2$
FGM	$u_1 u_2 (1 + \theta (1 - u_1)(1 - u_2))$
Gaussian	$\Phi_G[\Phi^{-1}(u_1), \Phi^{-1}(u_2); \theta]$
Clayton	$(u_1^{-\theta} + u_2^{-\theta} - 1)^{-1/\theta}$
Frank	$-\frac{1}{\theta}\log\left(1+\frac{(e^{-\theta u_1}-1)(e^{-\theta u_2}-1)}{e^{-\theta}-1}\right)$
Ali-Mikhail-Haq	$u_1 u_2 (1 - \theta (1 - u_1)(1 - u_2))]^{-1}$

• Relationship with Spearman's correlation coefficient and Kendall's correlation coefficient

Both $\rho_S(X,Y)$ and $\rho_\tau(X,Y)$ can be expressed in terms of copulas as follows:

$$\rho_S(X,Y) = 12 \int_0^1 \int_0^1 \{C(u_1, u_2) - u_1 u_2\} du_1 du_2,$$

$$\rho_\tau(X,Y) = 4 \int_0^1 \int_0^1 C(u_1, u_2) dC(u_1, u_2) - 1$$

• Relationship with Spearman's correlation coefficient and Kendall's correlation coefficient

Copula type	Function $C(u_1, u_2)$	θ -domain	Kendall's τ	Spearman's ρ
Product	$u_1 u_2$	N.A.	0	0
FGM	$u_1 u_2 (1 + \theta (1 - u_1)(1 - u_2))$	$-1 \leq \theta \leq +1$	$\frac{2}{9} heta$	$\frac{1}{3} heta$
Gaussian	$\Phi_G[\Phi^{-1}(u_1), \Phi^{-1}(u_2); \theta]$	$-1 < \theta < +1$	$\frac{2}{\pi} \arcsin(\theta)$	$\frac{6}{\pi} \arcsin(\frac{\theta}{2})$
Clayton	$(u_1^{-\theta} + u_2^{-\theta} - 1)^{-1/\theta}$	$\theta \in (0,\infty)$	$\frac{\theta}{\theta+2}$	*
Frank	$-\frac{1}{\theta} \log \left(1 + \frac{(e^{-\theta u_1} - 1)(e^{-\theta u_2} - 1)}{e^{-\theta} - 1} \right)$	$\theta \in (-\infty,\infty)$	$1 - \frac{4}{\theta} [1 - D_1(\theta)]$	$1 - \frac{12}{\theta} [D_1(\theta) - D_2(\theta)]$
Ali-Mikhail-Haq	$u_1 u_2 (1 - \theta (1 - u_1)(1 - u_2))]^{-1}$	$-1 \leq \theta \leq 1$	$\left(\frac{3\theta-2}{\theta}\right)$	*
			$-\frac{2}{3}(1-\frac{1}{\theta})^2\ln(1-\theta)$	

 $D_k(x)$ denotes the "Debye" function $k/x^k \int_0^x \frac{t^k}{(e^t-1)} dt, \ k=1,2$



* Some notes are from Climate Data Analysis course

Copulas and Dependence



• Goodness-of-fit test:

1) Graphical comparison: Theoretical vs. Empirical

2) Compute Maximum log-likelihood

3) p-value test

• Package: copula

Author: Marius Hofert, Ivan Kojadinovic, Martin Maechler, and Jun Yan

R code:

```
setwd("C:/Users/HRG/Desktop")
library(copula)
da90<-read.delim("marxy.txt",header=FALSE, sep="\t", dec=".")
names(da90)<- c("Prcp", "Temp")
attach(da90)
u<-pobs(da90[,1:2])</pre>
```

```
fc<-frankCopula(dim=2)
ffc<-fitCopula(fc,u)
```

```
nc<-normalCopula(dim=2, dispstr="un")
fnc<-fitCopula(nc,u)</pre>
```

fgc@loglik; fcc@loglik; ffc@loglik; fnc@loglik; ftc@loglik; fpc@loglik; fjc@loglik;

Example 1: CA 2014 temperature and precipitation



• AghaKouchak A., Cheng L., Mazdiyasni O., Farahmand A., 2014, Global Warming and Changes in Risk of Concurrent Climate Extremes: Insights from the 2014 California Drought, *Geophysical Research Letters*

Autocorrelation:



Goodness-of-fit test:

RMSE	GEV	Lognorm	GP	Gamma	Exp	Well
Temp	0.085	0.099	0.093	0.105	0.125	0.108
Pcpn	0.014	0.048	0.029	0.012	0.052	0.012

Goodness-of-fit test:



Marginal:



Goodness-of-fit test:

t copula fit (red curve) against empirical (black dashed lines)



Prcp Anomaly

Goodness-of-fit test:



0.8

4.0

0.0

0.0

0.2

0.

Volume



Peak

Peak

0.6

0.8

1.0

Goodness-of-fit test:

	Parameter	loglikelihood	p-value
Gumbel	NA	NA	NA
Clayton	0.198	1.53	0.6479
Frank	0.469	0.36	0.1833
Normal	0.096	0.42	0.2672
t	0.096	0.42	0.2772

Goodness-of-fit test:

	Parameter	loglikelihood	p-value
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Clayton	0.198	1.53	0.6479
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Normal	0.096	0.42	0.2672
t	0.096	0.42	0.2772

Example 1: **Bivariate Return Period**: joint analysis of temp and pcpn



Streamflow data:









t copula fit against empirical



sf1

What is the return period *T* of the univariate of CA 2014 precipitation?

 $T = \frac{m}{1 - p}$

where m > 0 is the average interarrival time of two consecutive events; p is the non-exceedance probability.



Example 2:

Bivariate Return Period: analysis of CA drought duration and severity



Time

Figure is from Mirabbasi 2012 paper

Example 2:

<u>Univariate</u>: the current CA drought duration is 3 years (ranked 7th) the 3-year precipitation deficit is 522 mm (ranked 3rd)

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<u>Univariate</u>: the current CA drought duration is 3 years (ranked 7th) the 3-year precipitation deficit is 522 mm (ranked 3rd)



Example 2: **Bivariate Return Period**: Joint analysis of CA drought duration and severity



 Cheng L., Hoerling M., AghaKouchak A., Livneh B., Quan X., 2015, Current Effects of Human-induced Climate Change on California Drought, *Journal of Climate* (in press)

Example 2: **<u>Bivariate Return Period</u>**: Joint analysis of CA drought duration and severity



Example 3: Precipitation and Soil Moisture (at 10cm) from preindustrial and industrial periods



More Applications

- 1. Medicine: Estimate the effect of an endogenous binary regressor (the "treatment") on a binary health outcome variable.
- 2. Finance: estimate the credit risk and the market risk.
- 3. Insurance
- 4. Biology
- 5. Health and environmental science
- 6. ...

- 1. Nonstationarity of the dependence structure (change-point)
- 2. Conditional predictability (ungauged point or time step)
- 3. Spatial dependence:
 - e.g. 10 out of 100 stations get flooding in a watershed. In a changing climate, the number of flooding stations increases to15. What are the plausible reasons?
- 4. Parameter uncertainty estimation (Bayesian inference)